

FOR $-\infty < F < \infty$

F IN Hz

↑ CAPITAL NOW

PSD IS AGAIN FOURIER TRANSFORM OF ACF

$$P_x(F) = \int_{-\infty}^{\infty} r_x(\tau) e^{-j2\pi F\tau} d\tau$$

PROPERTIES:

17.19 - PSD IS REAL FUNCTION

$$P_x(F) = \int_{-\infty}^{\infty} r_x(\tau) \cos(2\pi F\tau) d\tau$$

17.20 - $P_x(F) \geq 0$

17.21 - $P_x(-F) = P_x(F)$ (EVEN)

17.22 - ACF RECOVERED FROM
INVERSE FOURIER TRANSFORM

$$\begin{aligned} r_x(\tau) &= \int_{-\infty}^{\infty} P_x(F) e^{j2\pi F\tau} dF \\ &= \int_{-\infty}^{\infty} P_x(F) \cos(2\pi F\tau) dF \end{aligned}$$

NOTE: PSD IS NOT PERIODIC NOW.

CONTINUOUS-TIME WGN

AS A MODEL FOR PHYSICAL NOISE WE
ASSUME $x(t)$ IS GAUSSIAN WITH ZERO
MEAN FOR ALL t AND IS WSS WITH

PSD $P_x(f) = \frac{N_0}{2} \quad -\infty < f < \infty$

NOT PHYSICALLY POSSIBLE SINCE
 $r_x(0) = \int_{-\infty}^{\infty} P_x(f) df = \int_{-\infty}^{\infty} \frac{N_0}{2} df \rightarrow \infty$

HOWEVER, ANY SYSTEM WILL HAVE FINITE BANDWIDTH W SO THAT POWER WILL BE FINITE.

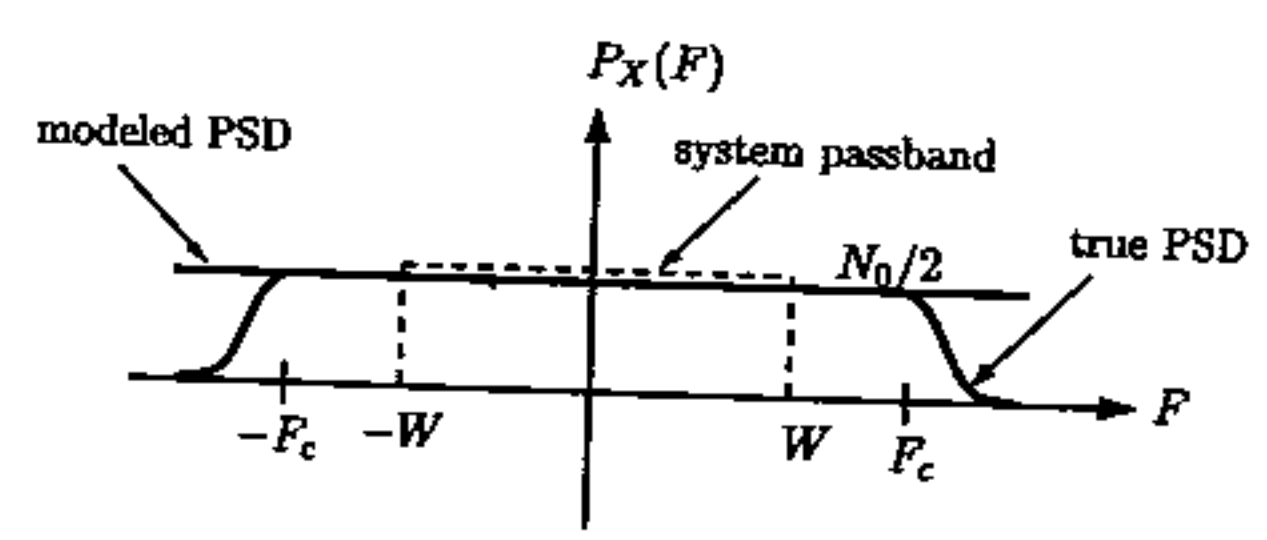
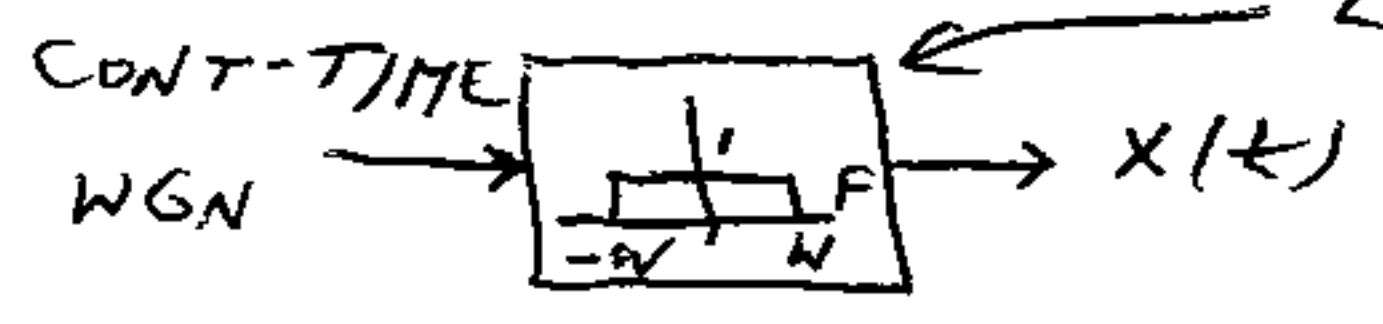


Figure 17.14: True and modeled PSDs for continuous-time white Gaussian noise.

NOTE THAT $r_x(\tau) = \int_{-\infty}^{\infty} P_x(f) e^{j2\pi f\tau} df$
 $= \int_{-\infty}^{\infty} \frac{N_0}{2} e^{j2\pi f\tau} df$
 $= \frac{N_0}{2} \delta(\tau)$

⇒ ALL SAMPLES ARE UNCORRELATED (EVEN IF VERY CLOSE TO EACH OTHER)

NOW CONSIDER OUTPUT OF SYSTEM WITH BANDWIDTH W LOWPASS FILTER



$$\Rightarrow P_x(f) = \begin{cases} \frac{N_0}{2} & |f| \leq W \\ 0 & |f| > W \end{cases}$$

IF WE USE NYQUIST SAMPLING OR
 $F_s = 2W$ SAMPLES/SEC (TO INPUT DATA
 INTO A COMPUTER)

$$x(t) \Big|_{t=n\Delta_t} = x[n] \quad -\infty < n < \infty$$

\uparrow
 $F_s = 1/\Delta_t$

$\Delta_t =$ TIME INTERVAL
BETWEEN SAMPLES

$x[n]$ = DISCRETE-TIME GAUSSIAN R.P.
 WITH ZERO MEAN AND ACS

$$\begin{aligned} r_x[k] &= E[x[n]x[n+k]] \\ &= E[x(n\Delta_t)x((n+k)\Delta_t)] \\ &= r_x(k\Delta_t) \leftarrow \text{SAMPLED VERSION} \\ &\quad \text{OF } r_x(\tau) \end{aligned}$$

$$\text{BUT } r_x(\tau) = \mathcal{F}^{-1} \left\{ \begin{array}{l} N_0/2 \quad |f| \leq W \\ 0 \quad |f| > W \end{array} \right\}$$

$$= \int_{-\infty}^{\infty} P_x(f) e^{j2\pi f\tau} df$$

$$= \int_{-W}^W \frac{N_0}{2} e^{j2\pi f\tau} df$$

$$= \int_{-W}^W \frac{N_0}{2} \cos(2\pi f\tau) df \quad \text{WHY?}$$

$$\begin{aligned}
 &= \frac{N_0}{2} \left. \frac{\sin 2\pi f \tau}{2\pi \tau} \right|_{-W}^W \\
 &= \frac{N_0}{2} \left[\frac{\sin(2\pi W \tau) - \sin(-2\pi W \tau)}{2\pi \tau} \right] \\
 &= \frac{N_0 \sin(2\pi W \tau)}{2\pi \tau} = N_0 W \frac{\sin(2\pi W \tau)}{2\pi W \tau}
 \end{aligned}$$

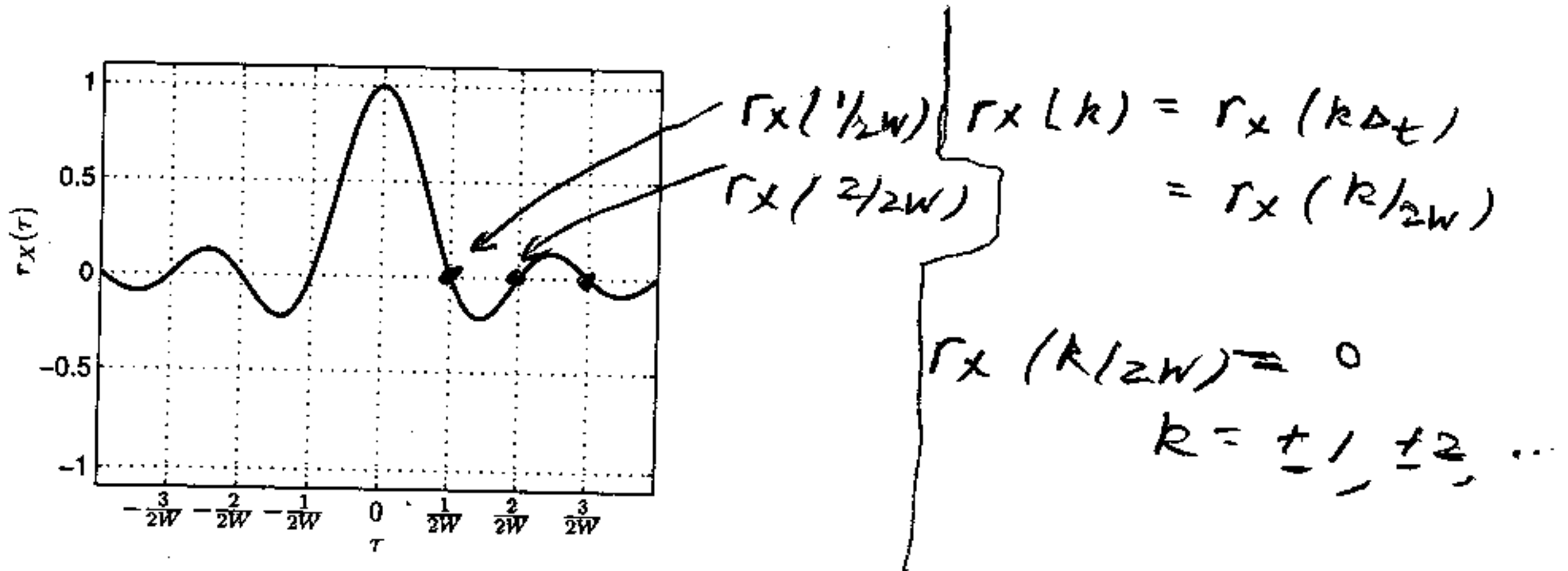
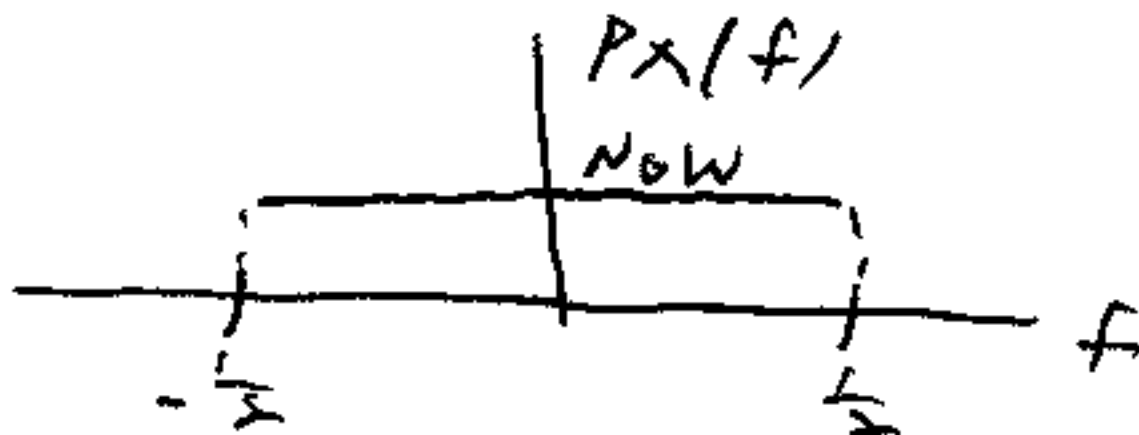


Figure 17.15: ACF for bandlimited continuous-time WGN with $N_0 W = 1$.

$$\begin{aligned}
 \Rightarrow r_x[k] &= 0 & k = \pm 1, \pm 2, \dots \\
 &= r_x[0] & k = 0 \\
 &= N_0 W & k = 0
 \end{aligned}$$

$$\Rightarrow r_x[k] = N_0 W \delta[k]$$

OR $X[n]$ IS DISCRETE-TIME WGN WITH VARIANCE $\sigma^2 = N_0 W$



JUSTIFIES USE OF WGN FOR SYSTEMS ANALYSIS IN DISCRETE-TIME

CHAPTER 18 - LINEAR SYSTEMS AND WSS RPS

CONSIDER ONLY DISCRETE-TIME SYSTEMS (READ APP. D). ASSUME WE INPUT A WSS R.P. $U[n]$ TO A LINEAR SHIFT INVARIANT (LSI) SYSTEM.

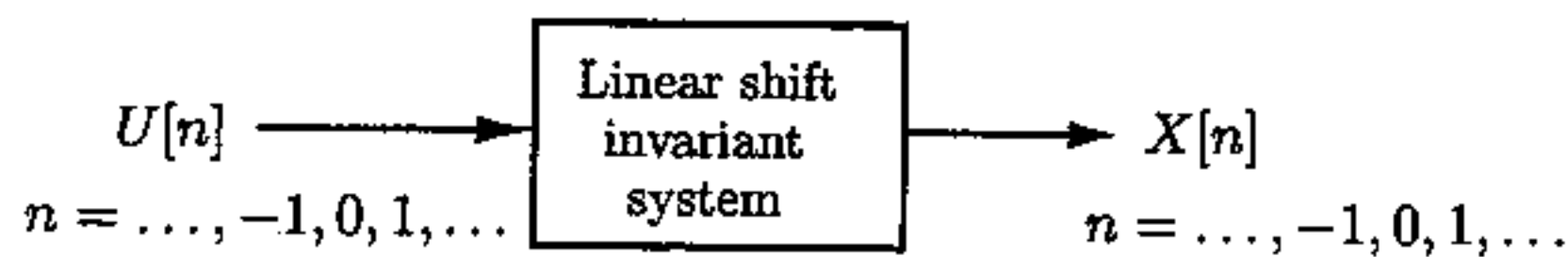


Figure 18.1: Linear shift invariant system with random process input and output.

WHAT CAN BE SAID ABOUT $x[n]$?

EXAMPLE: RECALL MA PROCESS (GENERAL)
 $x[n] = \frac{1}{2} (u[n] + u[n-1])$

↑ ZERO MEAN WSS
 $\Gamma_U[k] = \sigma_U^2 \delta[k]$
 (WHITE NOISE)

NO NEED TO ASSUME $U[n]$ IS GAUSSIAN
 IN THIS CHAPTER - INTERESTED IN ONLY
 FIRST TWO MOMENTS OF $x[n]$.

MA PROCESS CAN BE VIEWED AS
 OUTPUT OF LSI SYSTEM (= FILTER)

WITH IMPULSE RESPONSE

$$h[n] = \begin{cases} \frac{1}{2} & n=0 \\ \frac{1}{2} & n=1 \\ 0 & \text{OTHERWISE} \end{cases}$$

RECALL FOR LSI SYSTEM

$$x[n] = \sum_{k=-\infty}^{\infty} h[k] u[n-k] = h[n] \star u[n]$$

↑
CONVOLUTION

FOR MA RF,

$$\begin{aligned} x[n] &= h[0]u[n] + h[1]u[n-1] \\ &= \frac{1}{2}u[n] + \frac{1}{2}u[n-1] \\ &= \frac{1}{2}(u[n] + u[n-1]) \end{aligned}$$

SOME OTHER CHARACTERIZATIONS OF LSI SYSTEM:

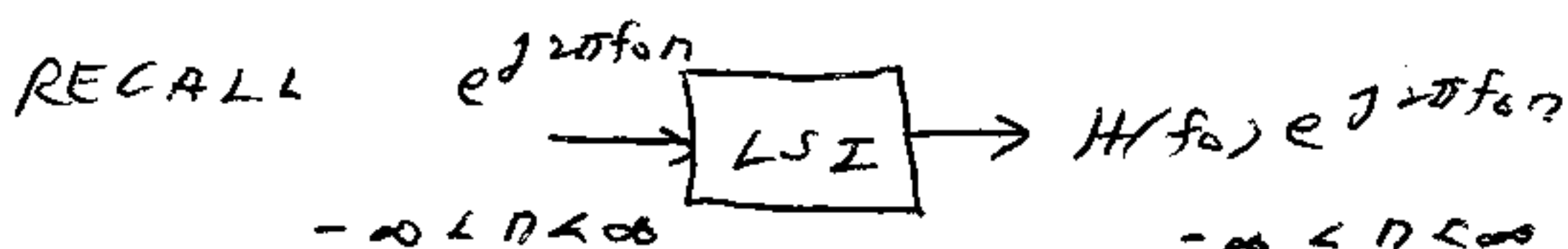
1) SYSTEM FUNCTION = Z-TRANSFORM OF IMPULSE RESPONSE

$$\begin{aligned} H(z) &= \sum_{k=-\infty}^{\infty} h[k] z^{-k} \\ &= \frac{1}{2}z^0 + \frac{1}{2}z^{-1} = \frac{1}{2} + \frac{1}{2}z^{-1} \quad \text{MA} \end{aligned}$$

2) FREQUENCY RESPONSE - FOURIER TRANSFORM OF IMPULSE RESPONSE

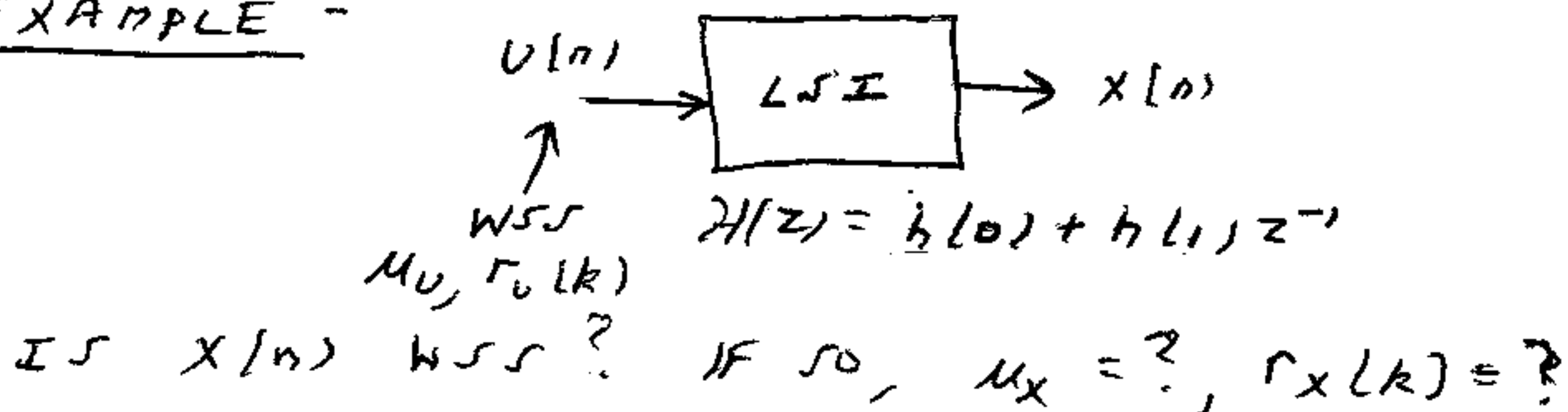
$$\begin{aligned}
 H(f) &= \sum_{k=-\infty}^{\infty} h[k] e^{-j2\pi f k} & |f| \leq \frac{1}{2} \\
 &= \frac{1}{2} e^{-j2\pi f(0)} + \frac{1}{2} e^{-j2\pi f(1)} \\
 &= \frac{1}{2} + \frac{1}{2} e^{-j2\pi f} & \text{MA}
 \end{aligned}$$

ALSO, $H(f) = \mathcal{H}(z) \Big|_{z=e^{j2\pi f}}$



$|H(f_0)| =$ MAGNITUDE RESPONSE AT $f=f_0$
 $\angle H(f_0) =$ PHASE RESPONSE AT $f=f_0$

EXAMPLE -



$$x[n] = h[0]v[n] + h[1]v[n-1] \quad -\infty < n < \infty$$

$$\begin{aligned}
 E[x[n]] &= h[0] \underbrace{E[v[n]]}_{\mu_v} + h[1] \underbrace{E[v[n-1]]}_{\mu_v} \\
 &= (h[0] + h[1]) \mu_v
 \end{aligned}$$

$E\{x(n)\}$ NOT DEPENDENT ON n

$$\mu_x = (h(0) + h(1))\mu_v$$

IN GENERAL, $\mu_x = \left. \sum_{k=-\infty}^{\infty} h(k) e^{-j2\pi f k} \right|_{f=0} \mu_v$

$$= H(0)\mu_v \quad \text{WHY?}$$

$$E\{x(n)x(n+k)\} = E\left\{ (h(0)v(n) + h(1)v(n-1)) \right. \\ \left. (h(0)v(n+k) + h(1)v(n+k-1)) \right\}$$

$$= h^2(0)E\{v(n)v(n+k)\} + h(0)h(1)E\{v(n)v(n+k-1)\} \\ + h(1)h(0)E\{v(n-1)v(n+k)\} + h^2(1)E\{v(n-1)v(n+k-1)\}$$

$$= h^2(0)r_v(k) + h(0)h(1)r_v(k-1) + h(1)h(0)r_v(k+1) \\ + h^2(1)r_v(k)$$

$$= (h^2(0) + h^2(1))r_v(k) + h(0)h(1)r_v(k-1) \\ + h(1)h(0)r_v(k+1)$$

$$= r_x(k) \quad \text{NOT DEPENDENT ON } n$$

$\Rightarrow x(n)$ IS WSS

ALSO, NOTE THAT IF

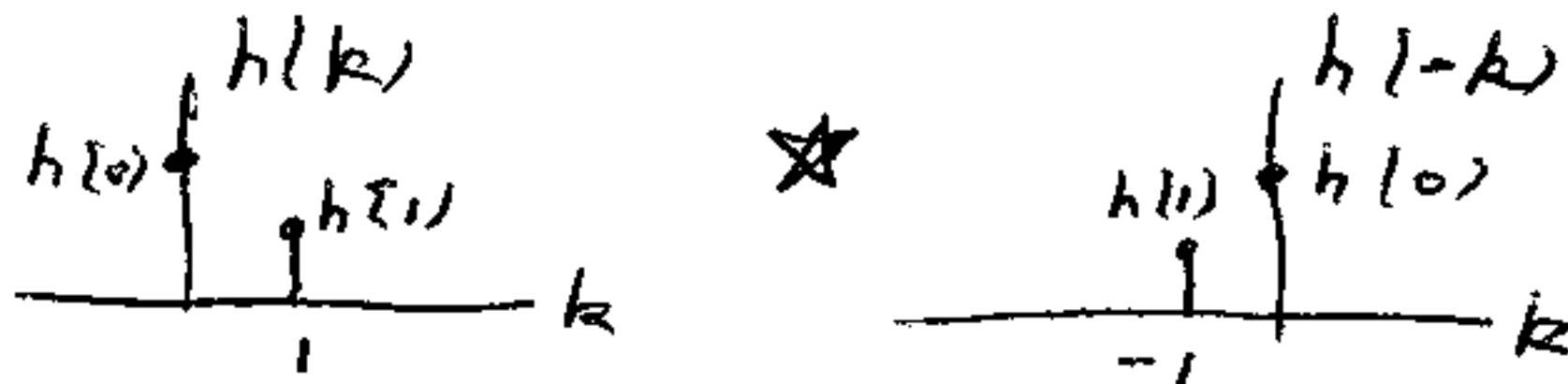
$$g(0) = h^2(0) + h^2(1)$$

$$g(1) = h(0)h(1)$$

$$g(-1) = h(1)h(0)$$

$$\begin{aligned}
 r_x(k) &= g[0]r_v(k) + g[1]r_v(k-1) + g[-1]r_v(k+1) \\
 &= \sum_{j=-1}^1 g[j]r_v(k-j) \\
 &= g[k] \star r_v(k)
 \end{aligned}$$

ALSO, $g[k] = h[k] \star h[-k]$

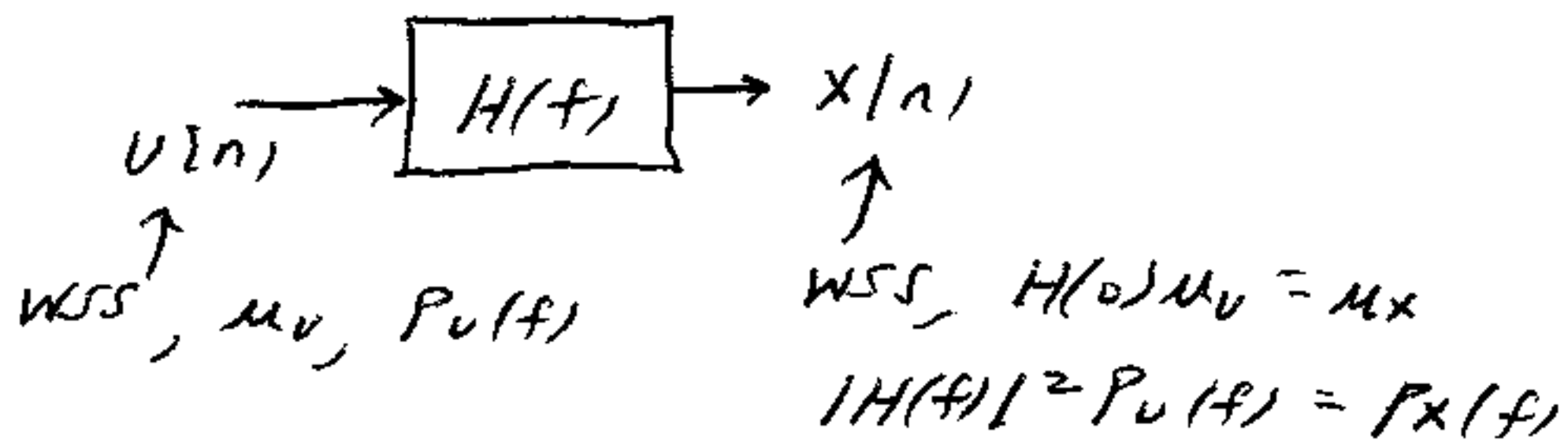


$$\begin{aligned}
 P_x(f) &= \mathcal{F}\{r_x(k)\} = \mathcal{F}\{g[k] \star r_v(k)\} \\
 &= G(f)P_v(f)
 \end{aligned}$$

BUT $\mathcal{F}\{h[k]\} = H(f)$ FREQUENCY RESPONSE
 $\mathcal{F}\{h[-k]\} = H^*(f)$

$$\begin{aligned}
 \Rightarrow P_x(f) &= G(f)P_v(f) \\
 &= \mathcal{F}\{h[k] \star h[-k]\} P_v(f) \\
 &= \mathcal{F}\{h[k]\} \mathcal{F}\{h[-k]\} P_v(f) \\
 &= H(f)H^*(f)P_v(f)
 \end{aligned}$$

$$\therefore P_x(f) = |H(f)|^2 P_v(f)$$



SEE GENERAL PROOF - THEOREM 18.3.1

SPECIAL CASE: IF INPUT IS WHITE NOISE, $P_U(f) = \sigma_v^2$

$$\Rightarrow P_X(f) = |H(f)|^2 P_U(f) \\ = |H(f)|^2 \sigma_v^2$$

FILTER "COLORS" THE NOISE - MAKES NOISE CORRELATED AT OUTPUT

EXAMPLE: AR P.P.

RECALL $X[n] = aX[n-1] + u[n]$

↑ WHITE NOISE

NEED $H(f)$. BUT $H(f) = \mathcal{H}(z) |_{z=e^{j2\pi f}}$

AND

$$\mathcal{H}(z) = \frac{X(z)}{U(z)}$$

IF $X[n] = aX[n-1] + u[n]$

$u[n], X[n]$
ARE SEQUENCES

$$X(z) = a z^{-1} X(z) + U(z)$$

$$\Rightarrow \mathcal{H}(z) = \frac{1}{1 - a z^{-1}}$$

$$H(f) = \frac{1}{1 - a e^{-j2\pi f}}$$

$$\begin{aligned}
 P_x(f) &= |H(f)|^2 P_v(f) \\
 &= \frac{1}{|1 - a e^{-j2\pi f}|^2} \sigma_v^2 \\
 &= \frac{\sigma_v^2}{|1 - a e^{-j2\pi f}|^2} \quad |f| \leq \frac{1}{2}
 \end{aligned}$$

SAME RESULT AS BEFORE. TO FIND ACS CAN TAKE INVERSE FOURIER TRANSFORM. EASIER APPROACH IS AS FOLLOWS:

$$\begin{aligned}
 P_x(f) &= H(f) H^*(f) P_v(f) \\
 &= H(f) H^*(f) \sigma_v^2 \\
 \Rightarrow r_x(k) &= (h[k] \star h[-k]) \sigma_v^2 \\
 &= \sum_{i=-\infty}^{\infty} h[-i] h[k-i] \sigma_v^2
 \end{aligned}$$

LET $m = -i$

$$= \sigma_v^2 \sum_{m=-\infty}^{\infty} h[m] h[m+k]$$

CORRELATION OF $h[k]$ WITH ITSELF

FOR AR RP. WE ASSUME LSI SYSTEM IS CAUSAL $\Rightarrow h[m] = 0 \quad m < 0$

$$\begin{aligned}
 r_x[k] &= \sigma_v^2 \sum_{m=0}^{\infty} h[m] h[m+k] & k \geq 0 \\
 r_x[-k] & & k < 0
 \end{aligned}$$

WHAT IS $h[m]$?

RECALL FROM SLIDE 160

$$x[n] = \sum_{l=0}^{\infty} a^l v[n-l]$$

(OR $x[n] = a x[n-1] + \delta[n]$ AND LET $x[-1] = 0$ WHY?, RECURSE FORWARD TO FIND $x[n] = h[n] = a^n \quad n \geq 0$)

$$\Rightarrow h[l] = \begin{cases} a^l & l \geq 0 \\ 0 & l < 0 \end{cases} = a^l u_s[l] \quad \begin{matrix} \uparrow \\ \text{UNIT} \\ \text{STEP} \end{matrix}$$

$$r_x[k] = \sigma_v^2 \sum_{m=0}^{\infty} a^m u_s[m] a^{m+k} u_s[m+k]$$

FOR $k \geq 0$ AND $m \geq 0 \quad u_s[m] = u_s[m+k] = 1$

$$r_x[k] = \sigma_v^2 \sum_{m=0}^{\infty} a^{2m+k} = \sigma_v^2 a^k \sum_{m=0}^{\infty} (a^2)^m$$

$$= \frac{\sigma_v^2}{1-a^2} a^k \quad k \geq 0$$

$$= \frac{\sigma_v^2}{1-a^2} a^{|k|} \quad -\infty < k < \infty$$

SEE ALSO MA EXAMPLE 18.3

INTERPRETATION OF PSD

CAN FINALLY PROVE THAT

$$\int_{f_1}^{f_2} P_x(f) df = \text{AVERAGE POWER IN } [f_1, f_2]$$

TO MEASURE AVERAGE POWER WE FILTER $X[n]$ BY LSI FILTER WITH A NARROW BAND AND DETERMINE POWER AT FILTER OUTPUT.

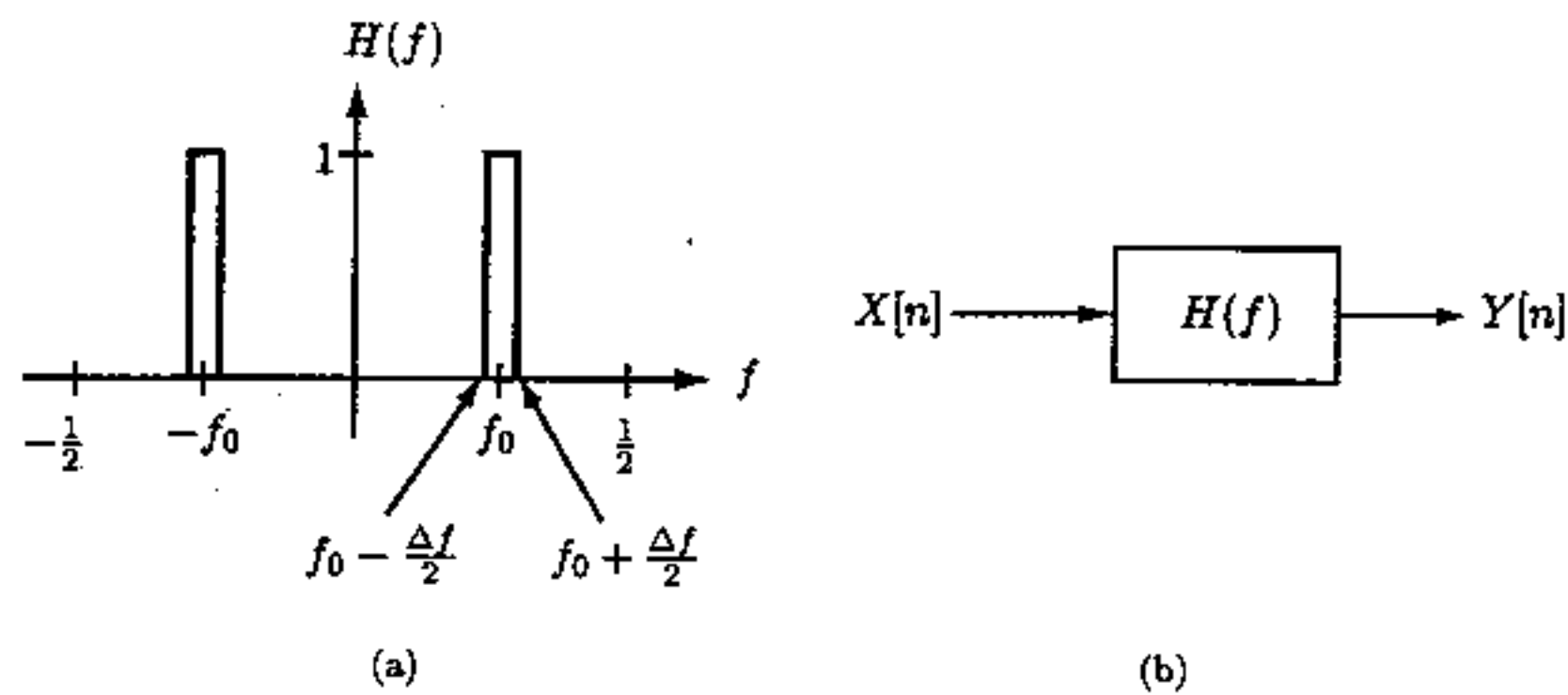


Figure 18.3: Narrowband filtering of random process to measure power within a band of frequencies.

$$\begin{aligned} \text{AVERAGE POWER OUT OF FILTER} &= E\{Y^2[n]\} \\ &= r_Y[0] \end{aligned}$$

$$\text{BUT } r_Y[k] = \int_{-\frac{1}{2}}^{\frac{1}{2}} P_Y(f) e^{j2\pi fk} df$$

$$\Rightarrow r_Y[0] = \int_{-\frac{1}{2}}^{\frac{1}{2}} P_Y(f) df$$

$$E\{Y^2[n]\} = \int_{-\frac{1}{2}}^{\frac{1}{2}} P_Y(f) df$$

$$= \int_{-\frac{1}{2}}^{\frac{1}{2}} |H(f)|^2 P_X(f) df$$

$$= \int_{f_0 - \frac{\Delta f}{2}}^{f_0 + \frac{\Delta f}{2}} 1 \cdot P_X(f) df$$

$$\neq \int_{f_0 - \frac{\Delta f}{2}}^{f_0 + \frac{\Delta f}{2}} 1 \cdot P_X(f) df$$

$$= 2 \int_{f_0 - \Delta f/2}^{f_0 + \Delta f/2} P_x(f) df \quad \text{WHY?}$$

TOTAL AVERAGE POWER IN BAND $[f_0 - \frac{\Delta f}{2}, f_0 + \frac{\Delta f}{2}]$

$$= \int_{f_0 - \Delta f/2}^{f_0 + \Delta f/2} P_x(f) df$$

VALID FOR
ANY Δf (NOT
NECESSARILY SMALL)

OR FOR $f_1 \leq f \leq f_2$

$$\text{TOTAL AVERAGE POWER} = \int_{f_1}^{f_2} P_x(f) df$$

(SIMILAR TO PDF!)

NOW LET $\Delta f \rightarrow 0$,

$$r_y(0) = 2 \int_{f_0 - \Delta f/2}^{f_0 + \Delta f/2} P_x(f) df$$

$$\rightarrow 2 P_x(f_0) \int_{f_0 - \Delta f/2}^{f_0 + \Delta f/2} df$$

$$= 2 P_x(f_0) \Delta f$$

$$P_x(f_0) = \frac{1}{2} \frac{r_y(0)}{\Delta f}$$

$$= \frac{1}{2} \frac{\text{TOTAL AVERAGE POWER OUT OF FILTER}}{\Delta f}$$

$$= \frac{1}{2} \frac{2 \times \text{TOTAL AVERAGE POWER IN } [f_0 - \Delta f/2, f_0 + \Delta f/2]}{\Delta f}$$

$$= \frac{\text{TOTAL AVERAGE POWER IN } [f_0 - \Delta f/2, f_0 + \Delta f/2]}{\Delta f}$$

= POWER PER UNIT FREQUENCY

= POWER SPECTRAL DENSITY !!

ORTHOGONALITY PRINCIPLE AGAIN

SEE SLIDES 92, 93 - WE NOW GENERALIZE THESE RESULTS. SEE PAGES 471-475 IN TEXT.

RECALL THAT FOR X, Y ZERO MEAN RVS WE CAN PREDICT Y BASED ON $X = x$ USING

$$\hat{y} = a_{\text{OPT}} x \quad \text{WHERE } a_{\text{OPT}} = \frac{\text{COV}(X, Y)}{\text{VAR}(X)}$$

THIS CHOICE OF a MINIMIZES THE MSE

$$\text{MSE}(a) = E[(Y - ax)^2]$$

OVER a . ALTERNATIVELY, a FOUND USING ORTHOGONALITY PRINCIPLE

$$E[(y - ax)x] = 0 \Rightarrow a_{opt} = \frac{E(xy)}{E(x^2)}$$

\uparrow ERROR \uparrow "DATA"

$$= \frac{COV(xy)}{VAR(x)}$$

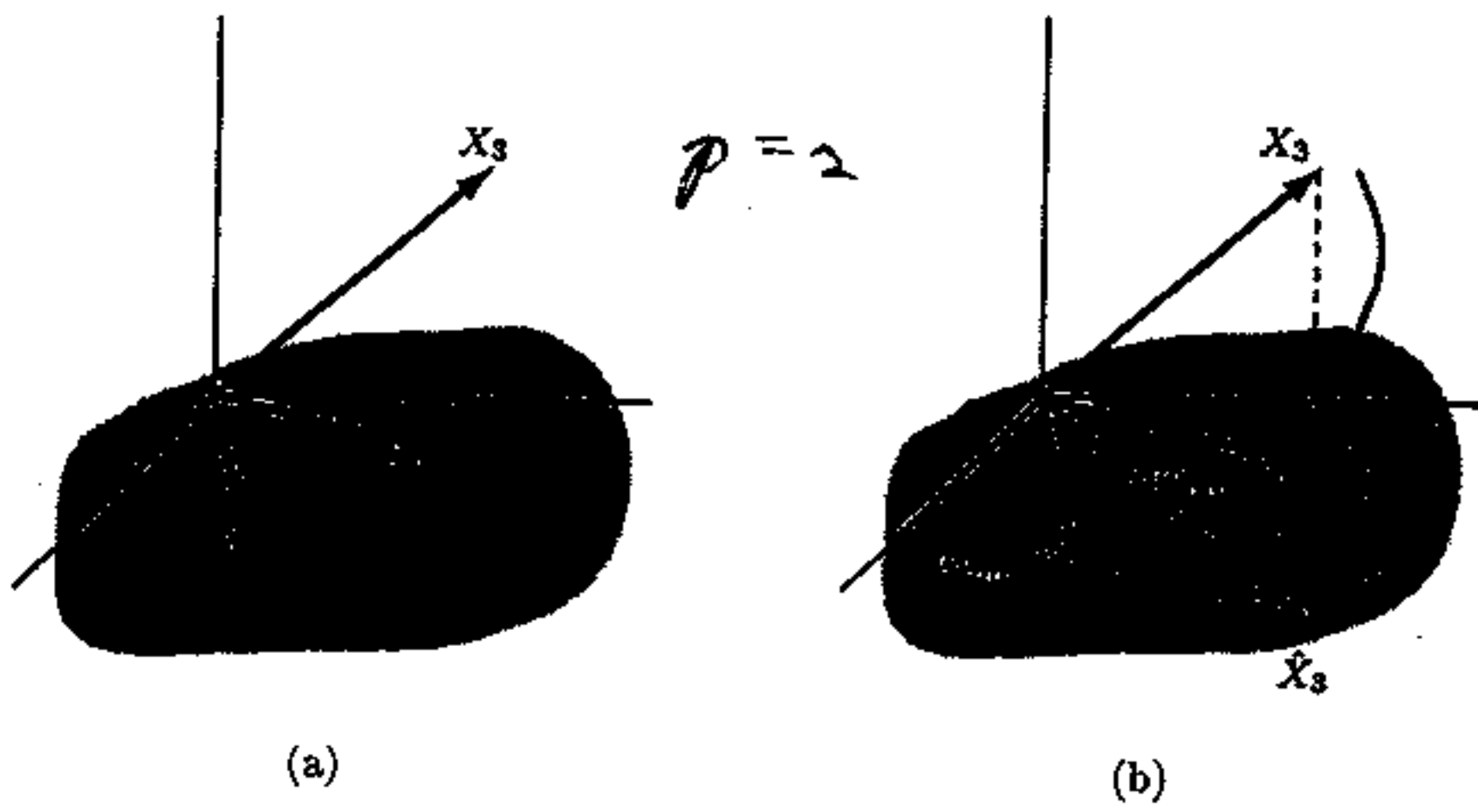
NOW ASSUME WE HAVE RVS

$X_1, X_2, \dots, X_p, X_{p+1}$ AND WE WISH TO PREDICT X_{p+1} BASED ON OBSERVING $x_1 = x_1, x_2 = x_2, \dots, x_p = x_p$ AS

$$\hat{X}_{p+1} = \sum_{i=1}^p a_i x_i$$

CHOOSE a_1, a_2, \dots, a_p TO MINIMIZE

$$MSE(a_1, a_2, \dots, a_p) = E_{x_1, \dots, x_{p+1}} \left[\left(X_{p+1} - \sum_{i=1}^p a_i x_i \right)^2 \right]$$



WANT $E \perp x_1$
 $E \perp x_2$
 $\Rightarrow E(\epsilon x_1) = 0$
 $E(\epsilon x_2) = 0$

GENERAL ORTHOGONALITY PRINCIPLE

Figure 14.3: Geometrical interpretation of linear prediction.

$$\Rightarrow E \left[\left(X_{p+1} - \sum_{i=1}^p a_i x_i \right) x_k \right] = 0 \quad k=1, 2, \dots, p$$

$$E [X_{p+1} X_k] = \sum_{i=1}^p a_i E [X_i X_k]$$

LET $C_{ij} = \text{COV}(X_i, X_j) = E \{ X_i X_j \} = C_{ji}$

$$\sum_{i=1}^p a_i C_{ik} = C_{p+1, k} \quad k=1, 2, \dots, p$$

OR $\sum_{i=1}^p C_{ki} a_i = C_{k, p+1} \quad k=1, 2, \dots, p$

SET OF p SIMULTANEOUS LINEAR EQUATIONS

$$\underbrace{\begin{bmatrix} C_{11} & C_{12} & \dots & C_{1p} \\ C_{21} & C_{22} & \dots & C_{2p} \\ \dots & \dots & \dots & \dots \\ C_{p1} & C_{p2} & \dots & C_{pp} \end{bmatrix}}_{\underline{C}} \underbrace{\begin{bmatrix} a_1 \\ a_2 \\ \vdots \\ a_p \end{bmatrix}}_{\underline{a}} = \underbrace{\begin{bmatrix} C_{1, p+1} \\ C_{2, p+1} \\ \vdots \\ C_{p, p+1} \end{bmatrix}}_{\underline{c}}$$

$$\underline{a}_{opt} = \underline{C}^{-1} \underline{c}$$

\underline{C} IS A COVARIANCE MATRIX \Rightarrow POSITIVE DEFINITE \Rightarrow INVERTIBLE

EXAMPLE : LET $p=2$

$$\underline{C}_x = \begin{bmatrix} 1 & 2/3 & 1/3 \\ 2/3 & 1 & 2/3 \\ 1/3 & 2/3 & 1 \end{bmatrix} = E [\underline{x} \underline{x}^T]$$

$$\underline{x} = \begin{bmatrix} X_1 \\ X_2 \\ X_3 \end{bmatrix}$$

PREDICT X_2 BASED ON $X_1 = x_1, X_2 = x_2$

$$\begin{bmatrix} c_{11} & c_{12} \\ c_{21} & c_{22} \end{bmatrix} \begin{pmatrix} a_1 \\ a_2 \end{pmatrix} = \begin{pmatrix} c_{13} \\ c_{23} \end{pmatrix}$$

$$\begin{bmatrix} 1 & 2/3 \\ 2/3 & 1 \end{bmatrix} \begin{pmatrix} a_1 \\ a_2 \end{pmatrix} = \begin{pmatrix} 1/3 \\ 2/3 \end{pmatrix}$$

$$\begin{pmatrix} a_1 \\ a_2 \end{pmatrix} = \frac{1}{1 - (2/3)^2} \begin{bmatrix} 1 & -2/3 \\ -2/3 & 1 \end{bmatrix} \begin{pmatrix} 1/3 \\ 2/3 \end{pmatrix}$$

$\underbrace{\hspace{10em}}_{\substack{\tilde{C}^T \\ \text{DET}(C)}}$ $\tilde{C} \equiv \text{COFACTOR}$
MATRIX

$$= \begin{pmatrix} -4/5 \\ 4/5 \end{pmatrix}$$

WHAT DO WE GET IF $C_x = \sigma^2 I$?

FINALLY ORTHOGONALITY PRINCIPLE HOLDS
IF $p \rightarrow \infty$.

PREDICT $X(n_0)$ BASED ON $\{X(n_0-1), X(n_0-2), \dots\}$
AS

$$\hat{X}(n_0) = \sum_{k=1}^{\infty} a_k X(n_0-k) \quad \text{AS AN EXAMPLE}$$

$$\Rightarrow E \left(\underbrace{(X(n_0) - \sum_{k=1}^{\infty} a_k X(n_0-k))}_{\text{ERROR}} \underbrace{X(n_0-l)}_{\substack{\uparrow \\ \text{"DATA"} \\ l=1, 2, \dots}} \right) = 0$$

WIENER FILTERING

WIENER FILTERING REFERS TO 4 SEPARATE BUT RELATED PROBLEMS

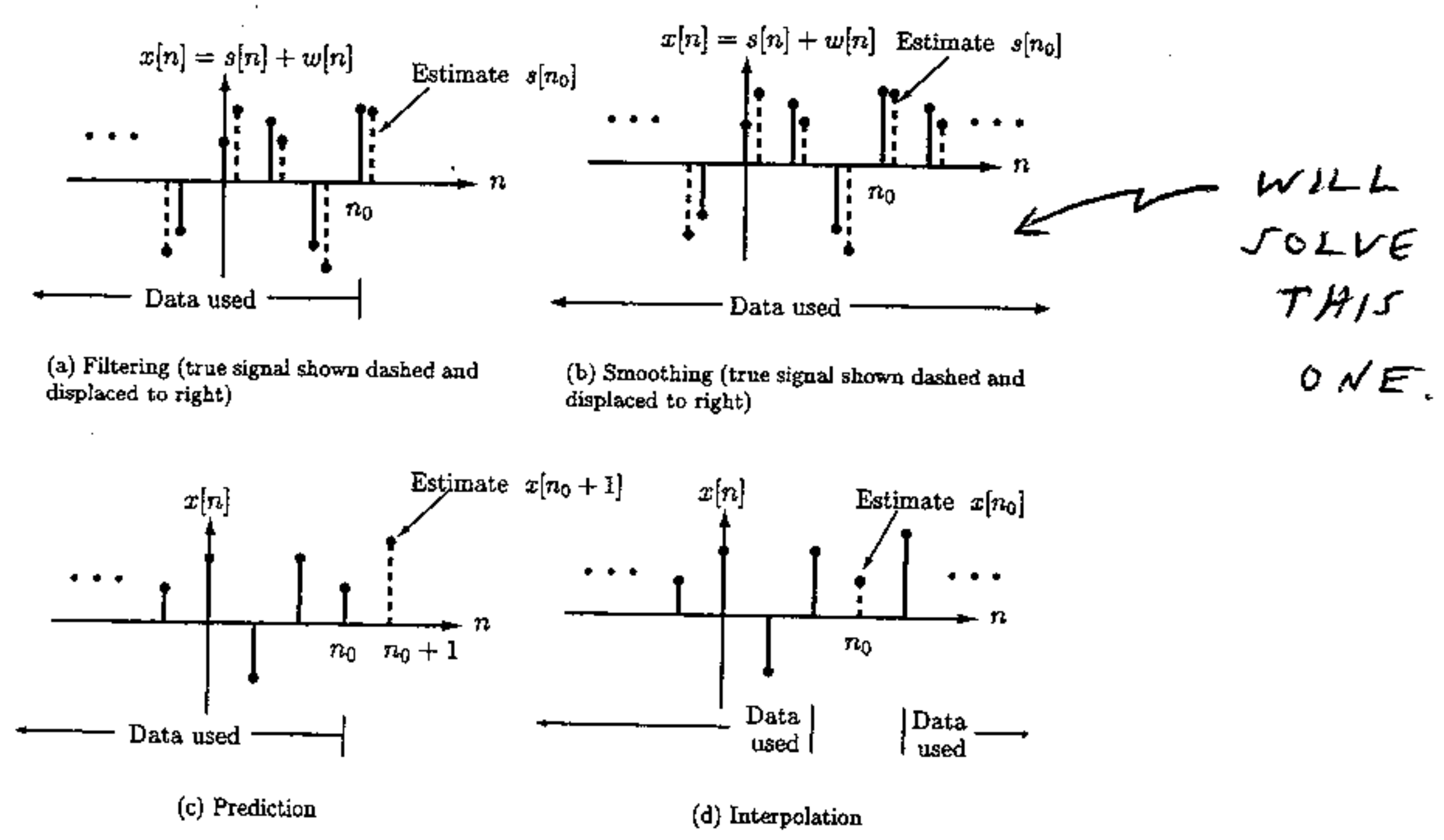


Figure 18.4: Definition of Wiener "filtering" problems.

SMOOTHING - NONREAL-TIME PROBLEM

WE OBSERVE $x[n] = s[n] + w[n] \quad -\infty < n < \infty$

ASSUMPTIONS: $s[n]$ IS ZERO MEAN WSS RP WITH PSD $P_s(f)$

$w[n]$ IS ZERO MEAN WSS RP WITH PSD $P_w(f)$

WISH TO ESTIMATE $s[n]$ FOR SOME $n = n_0$ USING A LINEAR FILTER

AS

$$\hat{S}[n_0] = \sum_{k=-\infty}^{\infty} h(k) X(n_0 - k)$$

↑ NONCAUSAL SINCE

$h(k) \neq 0$ FOR $k < 0$

ALSO, ESTIMATE AT $n = n_0$ BASED ON FUTURE SAMPLES $X(n_0 + 1), X(n_0 + 2), \dots$

↑ $h(-1)$ ↑ $h(-2)$

CALLED A WIENER SMOOTHER.

NOTE: JUST LIKE $\hat{X}_{p+1} = \sum_{i=1}^p a_i X_i$ BUT INFINITE NUMBER OF a_i 'S.

TO DETERMINE $h(k)$ WE MINIMIZE MSE OR

$$\text{MSE} = E \left[(S[n_0] - \hat{S}[n_0])^2 \right]$$

USING ORTHOGONALITY PRINCIPLE

$$E \left[(S[n_0] - \sum_{k=-\infty}^{\infty} h(k) X(n_0 - k)) X(n_0 - l) \right] = 0$$

$$E \left[S[n_0] X(n_0 - l) \right] = \sum_{k=-\infty}^{\infty} h(k) E \left[X(n_0 - k) X(n_0 - l) \right] \quad -\infty < l < \infty$$

BUT IT IS REASONABLE TO ASSUME THAT

$S(n)$ AND $W(n)$ ARE UNCORRELATED. SINCE THEY ARE ZERO MEAN $E \left[S(n) W(m) \right] = 0$ FOR ALL m, n .